

FM 5011 Fall 2011, Final Exam
Handout date: Thursday 15 December 2011

PRINT NAME:

SOLUTIONS

Remember to read to the bottom and to SIGN YOUR NAME BELOW!

Closed book, closed notes, no calculators/PDAs; no reference materials of any kind.
Show work; a correct answer, by itself, may be insufficient for credit.

I understand the above, and I understand that cheating has **severe consequences**, from a failing grade to expulsion.

REMEMBER TO SIGN YOUR NAME:

I. Definitions: Complete the following sentences.

a. (Topic 1500(28), 3 pts.) Let M be a set. A subset \mathcal{S} of 2^M is called a σ -algebra on M if ...

$$\left[\forall S_1, S_2, \dots \in \mathcal{A}, \quad S_1 \cup S_2 \cup \dots \in \mathcal{A} \right]$$
$$\& \left[\forall S \in \mathcal{A}, \quad M \setminus S \in \mathcal{A} \right]$$

b. (Topic 3000(37), 3 pts.) Let U be a random variable on the probability space $(\Omega, \mathcal{B}, \mu)$ and let \mathcal{S} be a σ -subalgebra of \mathcal{B} . A random variable X represents $E[U|\mathcal{S}]$ if ...

$$\left[X \text{ is } \mathcal{A}\text{-measurable} \right]$$
$$\& \left[\forall S \in \mathcal{A} \text{ of positive measure} \right. \\ \left. E[X|S] = E[U|S] \right]$$

c. (Topic 2900(20), 3 pts.) Let $X : \Omega \rightarrow \mathbb{R}$ be a random variable on $(\Omega, \mathcal{B}, \mu)$. The σ -algebra of X is $\mathcal{S}_X = \dots$

$$X^*(\mathcal{B}_{\mathbb{R}}), \quad \text{where } \mathcal{B}_{\mathbb{R}} \text{ is the} \\ \text{standard } \sigma\text{-algebra on } \mathbb{R}$$

d. (Topic 2330(23), 3 pts.) Let μ and ν be probability measures on a Borel space (Ω, \mathcal{B}) . We say that μ is **absolutely continuous** with respect to ν , and write $\mu \ll \nu$, if ...

$$\forall B \in \mathcal{B}, \quad (\nu(B) = 0) \implies (\mu(B) = 0)$$

e. (Topic 2700(28), 3 pts.) Let μ be a probability measure on \mathbb{R} and let λ denote Lebesgue measure on \mathbb{R} . A function $f : \mathbb{R} \rightarrow [0, \infty)$ is said to be a **probability density function** for μ if ...

$$\mu = f \lambda$$

f. (Topic 2900(8), 3 pts.) Let $X_{\bullet}^{(1)}, X_{\bullet}^{(2)}, \dots$ be a sequence of processes. Let X_{\bullet} be a process. We say $X_{\bullet}^{(n)}$ **converges to X_{\bullet} in finite dimensional marginals**, as $n \rightarrow \infty$, if ...

$$\forall \text{ integers } d \geq 1, \quad \forall t_1, \dots, t_d \geq 0,$$

$$\delta [X_{t_1}^{(k)}, \dots, X_{t_d}^{(k)}] \xrightarrow{k \rightarrow \infty} \delta [X_{t_1}^{(k)}, \dots, X_{t_d}^{(k)}]$$

g. (Topic 2900(23), 3 pts.) Two random variables X and Y are said to be **independent** if ...

$$\forall \text{ Borel } A, B \subseteq \mathbb{R},$$

$$(X \in A) \text{ is independent of } (Y \in B)$$

h. (Topic 2900(35), 3 pts.) Let μ and ν be two probability measures on \mathbb{R} . Define $A : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$ by $A(x, y) = x + y$. Then the **convolution** of μ and ν is given by $\mu * \nu = \dots$

$$A_* (\mu \times \nu)$$

II. True or False. (No partial credit.)

a. (Topic 2700(47), 2 pts.) Let X and Y be random variables. If their distributions have the same Fourier transforms, then $X = Y$ a.s.

F

b. (Topic 2900(9), 2 pts.) If X and Y are identically distributed random variables, then X^2 and Y^2 are also identically distributed.

T

c. (Topic 3600(21), 2 pts.) If V_\bullet and W_\bullet are Brownian motions, then $V = W$ in finite dimensional marginals.

T

d. (Topic 3000(37), 2 pts.) Let X and Y be random variables. Assume that X and Y are independent. Then $E[X|Y] = X$.

F

e. (Topic 2900(22), 2 pts.) If X is a random variable and $\mathcal{S} = \mathcal{S}_X$ is its σ -algebra, then X is \mathcal{S} -measurable.

T

f. (Topic 2900(24), 2 pts.) If X and Y are random variables, then $\delta_{X,Y} = \delta_X \times \delta_Y$.

F

g. (Topic 2900(44), 2 pts.) Let G be the grade of a standard normal random variable and let $0 \leq a < b \leq 1$. Then $\Pr[a < G < b] = b - a$.

T

h. (Topic 2400(7-8), 2 pts.) For any random variable X , there exists $a \in \mathbb{R}$, such that $\Pr[X < a] = 0$.

F

THIS PAGE IS FOR TOTALING SCORES
PLEASE DO NOT WRITE ON THIS PAGE

I. a-d.

I. e-h.

II. a-d.

II. e-h.

III(1).

III(2).

III(3,4).

III(5).

III(6).

III(7).

III(8).

III(9).

III(10).

III. Computations. Some of your answers may involve Φ , the cumulative distribution function of the standard normal distribution. (Unless otherwise specified, answers must be exactly correct, but can be left in any form easily calculated on a standard calculator.)

1. (Topic 3200(14-23), 20 pts.) Let X be a random variable whose distribution is χ^2 with two degrees of freedom. Compute $E[X^2]$.

Let Z_1, Z_2 be indep. std normal RVs

$$X \stackrel{\text{def}}{=} Z_1^2 + Z_2^2$$

$$\begin{aligned} X^2 &\stackrel{\text{def}}{=} (Z_1^2 + Z_2^2)^2 \\ &= Z_1^4 + 2Z_1^2 Z_2^2 + Z_2^4 \end{aligned}$$

Z_1^2, Z_2^2 indep & \therefore uncorrelated

$$E[X^2] = (E[Z_1^4]) + 2(E[Z_1^2])(E[Z_2^2]) + (E[Z_2^4])$$

$$= (3 \cdot 1) + 2(1)(1) + (3 \cdot 1)$$

$$= 3 + 2 + 3 = 8$$

2. (Topic 2900(47), 20 pts.) Let X_1, X_2, \dots be iid standard random variables. For all integers $n \geq 1$, let $Z_n := (X_1 + \dots + X_n)/\sqrt{n}$. Compute $\lim_{n \rightarrow \infty} E[(e^{Z_n} - e)_+]$

Z std normal RV

$$\| E[(e^Z - e)_+] \|$$

$$\| \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} (e^x - e)_+ e^{-x^2/2} dx \|$$

$$\| \frac{1}{\sqrt{2\pi}} \int_1^{\infty} (e^x - e) e^{-x^2/2} dx \|$$

$$e^{1/2} \left[\frac{1}{\sqrt{2\pi}} \int_0^{\infty} e^x e^{-x^2/2} dx \right] - e \left[\frac{1}{\sqrt{2\pi}} \int_1^{\infty} e \cdot e^{-x^2/2} dx \right]$$

$\|$

$$e^{1/2} [\Phi(0)] - e [\Phi(-1)]$$

3. (Topic 2900(13), 20 pts.) Let Z be a standard normal random variable. Let $\mu := \delta[Z^2]$ be the distribution of Z^2 . Compute $\int_{-\infty}^{\infty} x^5 d\mu(x)$.

$$E[(Z^2)^5]$$

//

$$E[Z^{10}]$$

//

$$1 \cdot 3 \cdot 5 \cdot 7 \cdot 9$$

distribution
drives
expectation

4. (Topic 2800(13), 20 pts.) Let $g(x) = x^5$. Let $\nu(x) = \begin{cases} x^2 + 2, & \text{if } x < 1; \\ x^4, & \text{if } x \geq 1. \end{cases}$

Compute $\int_0^2 g(x) d\nu(x)$.

$$\left[\int_0^1 x^5 (2x) dx \right] + \left[\int_1^1 (x^5) (-2) d\delta_1(x) \right] + \left[\int_1^2 x^5 (4x^3) dx \right]$$

$$2 \left[\int_0^1 x^6 dx \right] - 2 \left[\int_1^1 x^5 d\delta_1(x) \right] + 4 \left[\int_1^2 x^8 dx \right]$$

$$2 \left[\frac{x^7}{7} \right]_{x \rightarrow 0}^{x \rightarrow 1} - 2 \left[x^5 \right]_{x \rightarrow 1} + 4 \left[\frac{x^9}{9} \right]_{x \rightarrow 1}^{x \rightarrow 2}$$

$$2 \left[\frac{1}{7} \right] - 2 [1] + 4 \left[\frac{2^9 - 1}{9} \right]$$

5. (Topic 3200(2), 15 pts.) Assume that the distribution $\delta[X]$ of the random variable X has probability density function given by $p(x) = \frac{1}{\pi(1+x^2)}$. Let $Y := e^X$. Compute a probability density function f for $\delta[Y]$. Express $f(x)$ explicitly.

$$P := \text{CDF}_{\delta[X]}$$

$$F := \text{CDF}_{\delta[Y]}$$

$$P' = p$$

$$F' = f$$

$$\begin{aligned} F(x) &= P_n [Y \leq x] \\ &= P_n [e^X \leq x] \\ &= P_n [X \leq \ln x] \\ &= P(\ln x) \end{aligned}$$

$$\begin{aligned} f(x) = F'(x) &= [P'(\ln x)] \left[\frac{1}{x} \right] \\ &= [p(\ln x)] \left[\frac{1}{x} \right] \\ &= \left[\frac{1}{\pi(1+(\ln x)^2)} \right] \left[\frac{1}{x} \right] \end{aligned}$$

6. (Topic 3000(52), 15 pts.) Let C_1, C_2, C_3, \dots be iid binary random variables such that, for all integers $j \geq 0$, we have $\Pr[C_j = 1] = 0.5 = \Pr[C_j = -1]$. Let

$$X := E[(C_1 + \dots + C_{100}) \mid (C_1 + \dots + C_{50})].$$

Compute $E[X^2]$.

$$X = E[(C_1 + \dots + C_{50}) \mid (C_1 + \dots + C_{50})]$$

$$+ E[(C_{51} + \dots + C_{100}) \mid (C_1 + \dots + C_{50})]$$

$$= C_1 + \dots + C_{50}$$

$$+ E[C_1 + \dots + C_{50}]$$

$$= C_1 + \dots + C_{50}$$

$$X^2 = \sum_{j=1}^{50} \sum_{k=1}^{50} C_j C_k$$

$$E[X^2] = \sum_{j=1}^{50} \sum_{k=1}^{50} \delta_j^k = \sum_{j=1}^{50} 1 = 50$$

7. (Topic 3400(20), 15 pts.) Let X_1, \dots, X_{100} be iid normal variables with unknown mean μ and known variance 0.49. Let x_1, \dots, x_{100} be a sample modeled on X_1, \dots, X_{100} . Assume that the sample mean $(x_1 + \dots + x_{100})/(100) = 5$. Find a 99% confidence interval for μ . (Note: For a standard normal random variable Z , we have $\Pr[|Z| < 2.58] = 0.99$.)

$$\bar{X} := \frac{1}{100} [X_1 + \dots + X_{100}] \quad \bar{x} := \frac{1}{100} [x_1 + \dots + x_{100}] = 5$$

$$\sigma = \text{SD}[X_1] = \dots = \text{SD}[X_{100}] = \sqrt{0.49} = 0.7$$

$$E[\bar{X}] = \frac{1}{100} \left[\sum_{j=1}^{100} E[X_j] \right] = \frac{1}{100} [100\mu] = \mu$$

$$\text{Var}[\bar{X}] = \frac{1}{100^2} \left[\sum_{j=1}^{100} \text{Var}[X_j] \right] = \frac{1}{100^2} [100\sigma^2] = \frac{\sigma^2}{100}$$

$\frac{\bar{X} - \mu}{\sigma/10}$ is std normal

$$\Pr \left[\left| \frac{\bar{X} - \mu}{\sigma/10} \right| < 2.58 \right] = 0.99$$

With 99% confidence, $\left| \frac{\bar{x} - \mu}{\sigma/10} \right| < 2.58$,

i.e. $|\mu - \bar{x}| < (2.58) \left(\frac{\sigma}{10} \right)$

i.e. $|\mu - 5| < (2.58)(0.07)$

99% confidence interval:

$$\left(5 - (2.58)(0.07), \quad 5 + (2.58)(0.07) \right)$$

8. (Topic 3600(21), 15 pts.) Let W_\bullet be a Brownian motion. Compute $E[(W_4)^2(W_{13})^2]$.

$$Z_1 := \frac{W_4}{2}$$

$$2Z_1 = W_4$$

$$Z_2 := \frac{W_{13} - W_4}{3}$$

$$3Z_2 = W_{13} - W_4$$

$$2Z_1 + 3Z_2 = W_{13}$$

Z_1, Z_2 indep
std normal

$$E[(2Z_1)^2(2Z_1 + 3Z_2)^2]$$

||

$$E[4Z_1^2(4Z_1^2 + 12Z_1Z_2 + 9Z_2^2)]$$

||

$$E[16Z_1^4 + 48Z_1^3Z_2 + 36Z_1^2Z_2^2]$$

→ ||

$$16(E[Z_1^4]) +$$

$$48(E[Z_1^3])(E[Z_2]) +$$

$$36(E[Z_1^2])(E[Z_2^2])$$

→ ||

$$16(1 \cdot 3) +$$

$$48(0)(0) +$$

$$36(1)(1)$$

||

$$48 +$$

$$0 +$$

$$36$$

||

$$84$$

9. (Topic 3800(47), 10 pts.) Let W_t be a Brownian motion. Let X_t satisfy

$$dX_t/X_t = 2dW_t - 3dt, \quad X_0 = 1.$$

Compute $E[(X_4)^3]$.

Log-Ito

$$\begin{aligned} d(\ln X_t) &\stackrel{\text{Log-Ito}}{=} (2dW_t - 3dt) - \frac{1}{2}(2dW_t - 3dt)^2 \\ &= 2dW_t - 3dt - \frac{1}{2}(4dt) \\ &= 2dW_t - 5dt \end{aligned}$$

$$\begin{aligned} \ln X_4 &= (\ln X_0) + \int_0^4 (2dW_t - 5dt) \\ &= (\ln 1) + [2W_t - 5t]_{t=0}^{t=4} \\ &= (0) + [(2W_4 - 5 \cdot 4) - 0] \end{aligned}$$

Let Z be a
std normal RV

$$= 2W_4 - 20$$

$$\stackrel{\delta}{=} 4Z - 20$$

$$X_4 \stackrel{\delta}{=} e^{4Z - 20}$$

$$(X_4)^3 \stackrel{\delta}{=} e^{12Z - 60}$$

$$\begin{aligned} E[(X_4)^3] &= E[e^{12Z - 60}] = (e^{-60})(E[e^{12Z}]) \\ &= (e^{-60})(e^{12^2/2}) \end{aligned}$$

10. (Topic 0026, 10 pts.) Let X_t satisfy $dX_t = 4t dW_t + \frac{dt}{t^4 + 1}$.

Let Y_t be defined by $Y_t = t^3 e^{X_t}$. Compute $E \left[\int_0^4 \frac{dY_t}{e^{X_t}} \right]$.

$$dY_t = 3t^2 e^{X_t} dt + t^3 \left(e^{X_t} dX_t + \frac{1}{2} e^{X_t} (dX_t)^2 \right)$$

$$\frac{dY_t}{e^{X_t}} = 3t^2 dt + t^3 \left(\left(4t dW_t + \frac{dt}{t^4 + 1} \right) + \frac{1}{2} (16t^2 dt) \right)$$

$$= 4t^4 dW_t + \left(3t^2 + \frac{t^3}{t^4 + 1} + 8t^5 \right) dt$$

$$E \left[\int_0^4 \frac{dY_t}{e^{X_t}} \right] = \int_0^4 \left(3t^2 + \frac{t^3}{t^4 + 1} + 8t^5 \right) dt$$

$$= \left[t^3 + \frac{1}{4} (\ln(t^4 + 1)) + 8 \cdot \frac{t^6}{6} \right]_{t \rightarrow 0}^{t \rightarrow 4}$$

$$= 4^3 + \frac{1}{4} (\ln(4^4 + 1)) + 8 \cdot \frac{4^6}{6}$$