

# Manifold Regression via Brownian Motion

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August 19, 2015

# Kernel regression

## Nadaraya-Watson kernel estimator

Given observations  $\{(t_i, x_i)\}_{i=1}^n \in \mathbb{R} \times \mathbb{R}$ , learn the function  $f_0 : t \in \mathbb{R} \rightarrow x \in \mathbb{R}$ . The Nadaraya-Watson estimator is

$$\begin{aligned}\hat{f}_0(t) &= \frac{\sum_{i=1}^n K_h(t - t_i) x_i}{\sum_{i=1}^n K_h(t - t_i)} \\ &= \frac{\sum_{i=1}^n K_h(t - t_i) f_0(t_i)}{\sum_{i=1}^n K_h(t - t_i)} + \frac{\sum_{i=1}^n K_h(t - t_i) (x_i - f_0(t_i))}{\sum_{i=1}^n K_h(t - t_i)} \\ &\approx (I - h\Delta_{p^2(t)}) f_0 + (\text{error}) \\ &= f_0(t)\end{aligned}$$

when  $K_h$  is the Gaussian kernel.

Problem: One can not add  $x_i$  if they lie on a general manifold.

# Bayesian framework

## Basic idea

Bayes theorem:

$$\mathbb{P}(func|data) = \frac{\mathbb{P}(data|func) \cdot \mathbb{P}(func)}{\mathbb{P}(data)}$$

Terminology: Prior distribution :  $\mathbb{P}(func)$

Posterior distribution :  $\mathbb{P}(func|data)$

MAP estimator:  $\hat{f} = \operatorname{argmax}_{f \in C([0,1],M)} \mathbb{P}(f | \{(t_i, x_i)\}_{i=1}^n)$

Question: Consistency of this estimator requires the convergence of  $\hat{f}$  to  $f_0$ .

# Manifold regression

## Model set-up

Given a compact manifold  $M$  (e.g.,  $\mathbb{S}^D$ ,  $\text{SO}(3)$ , space of shapes) and observations  $\{(t_i, x_i)\}_{i=1}^n \in [0, 1] \times M$ , learn the function  $f_0 : [0, 1] \rightarrow M$ .

$$x|t \sim p_{\sigma^2}(f_0(t), x).$$

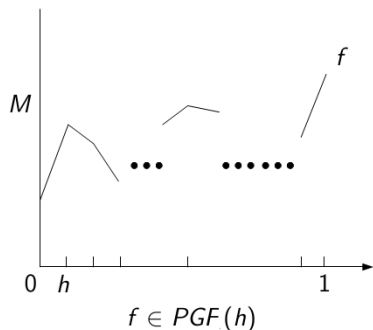
Here  $p_{\sigma^2}(f_0(t), x)$  denotes the heat kernel on  $M$ , determining how  $x_i$  deviates from  $f_0(t_i)$

When  $M = \mathbb{R}^D$ :

$$p_{\sigma^2}(f_0(t), \mathbf{x}) = \frac{1}{(\sqrt{2\pi}\sigma)^D} \exp\left(\frac{-\|\mathbf{x} - f_0(t)\|^2}{2\sigma^2}\right),$$

# Bayesian framework

Prior distribution (discrete BM)



The class of piecewise geodesic functions  $PGF(h) \subset C([0, 1], M)$

The prior distribution  $\Pi_{c,h}$  with density:

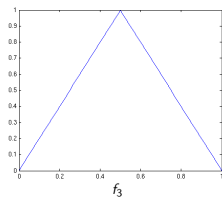
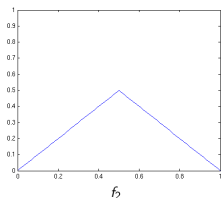
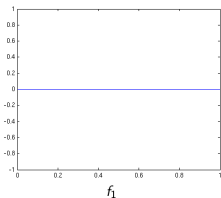
$$\pi_{c,h}(f) = \frac{1}{\mu(M)} \prod_{k=1}^{1/h} p_{ch}(f(kh-h), f(kh)).$$

# Bayesian framework

Prior distribution (discrete BM)

Let  $c = 1, h = 0.5$ ,  $f \in PGF(h) \subset C([0, 1], \mathbb{S}^1)$  is determined by 3 values  $f(0), f(0.5), f(1) \in \mathbb{S}^1 \equiv [0, 2\pi] / \{0, 2\pi\}$ .

	$f(0)$	$f(0.5)$	$f(1)$	$\pi_{c,h}(f)$
$f_1$	0	0	0	0.051
$f_2$	0	0.5	0	0.031
$f_3$	0	1	0	0.007



# Bayesian framework

## Prior distribution (continuous BM)

The continuous BM prior on  $\mathcal{P} \equiv C([0, 1], M)$  is the limit of the discrete BM prior. It is the unique probability measure  $\Pi$  such that for any  $n \in \mathbb{N}$ ,  $0 = t_0 < \dots < t_n = 1$ , and open subsets  $U_0, U_1, \dots, U_n \in M$ , the following identity is satisfied

$$\Pi(f \in \mathcal{P} \mid f(t_0) \in U_0, \dots, f(t_n) \in U_n) = \int_{U_0} \frac{d\mu(x_0)}{\mu(M)} \int_{U_1 \times \dots \times U_n} \prod_{i=1}^n p_{c(t_i - t_{i-1})}(x_i, x_{i-1}) d\mu(x_i).$$

# Bayesian framework

## Posterior distribution

Given observations  $\{(t_i, x_i)\}_{i=1}^n$  i.i.d drawn from

$$x|t \sim p_{\sigma^2}(f_0(t), x), \quad t \sim p(t)$$

the posterior distribution of  $\Pi$  has the density function

$$\begin{aligned}\Pi(f \in A | \{(t_i, x_i)\}_{i=1}^n) &\propto \int_{f \in A} \prod_{i=1}^n p(t_i, x_i | f) d\Pi(f) \\ &= \int_{f \in A} \prod_{i=1}^n p_{\sigma^2}(f(t_i), x_i) p(t_i) d\Pi(f).\end{aligned}$$

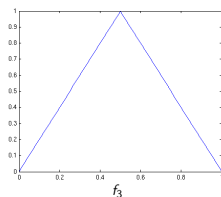
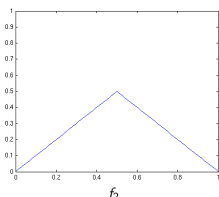
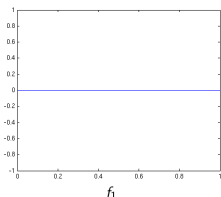


# Bayesian framework

## Posterior distribution

If  $\sigma^2 = 0.3$  and we observe  $(0.5, 0.6) \in [0, 1] \in \mathbb{S}^1$ , then

	$f(0)$	$f(0.5)$	$f(1)$	$\pi_{c,h}(f)$	$\pi_{c,h}(f f(0.5) = 0.6)$
$f_1$	0	0	0	0.051	0.020
$f_2$	0	0.5	0	0.031	0.022
$f_3$	0	1	0	0.007	0.004



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**Algorithm 1** Bayesian Manifold regression via discrete BM

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**Input:** Observations  $\{(t_i, x_i)\}_{i=1}^n \subset [0, 1] \times M$ , grid spacing  $h$  ( $m = 1/h + 1$ ), scaling constant  $c$ , variance  $\sigma^2$ , temperature  $T$  for Simulated Annealing (SA).

**Output:** A function in  $PGF(h)$  determined by  $\{(t_j^*, x_j^*)\}_{j=1}^m$ .

**Steps:**

- $t_j^* = (j - 1)h$ ,  $x_j^{(0)}$  i.i.d. uniformly sampled from  $M$  and  $f^{(0)} \in PGF(h)$  determined by  $\{(t_j^*, x_j^{(0)})\}_{j=1}^m$
- Energy function  $E(f) = \pi_{c,h}(f) \prod_{i=1}^n p_{\sigma^2}(f(t_i), y_i) p(t_i)$
- $f^* = SA(E, f^{(0)}, T)$

**return** The function  $f^*$  in  $PGF(h)$  determined by  $\{(t_j^*, x_j^*)\}_{j=1}^m$ .

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**Algorithm 2** Bayesian Manifold regression via continuous BM

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**Input:** Observations  $\{(t_i, x_i)\}_{i=1}^n \subset [0, 1] \times M$ ,  $\{t_j^*\}_{j=1}^m$ , scaling constant  $c$ , variance  $\sigma^2$ , temperature  $T$  for Simulated Annealing (SA).

**Output:** Predictions  $\{(t_j^*, x_j^*)\}_{j=1}^m$ .

**Steps:**

- Let  $\{t'_j\}_{i=1}^{n+m}$  be sorted array of  $\{t_i\}_{i=1}^n \cup \{t_j^*\}_{j=1}^m$ ,  $x_j^{(0)}$  i.i.d. uniformly sampled from  $M$  and  $f^{(0)}$  such that  $f^{(0)}(t'_j) = x_j^{(0)}$
- $\tilde{\pi}_c(f) = \prod_{i=1}^{n+m} p_{c(t'_i - t'_{i-1})}(f(t'_i), f(t'_{i-1}))$
- Energy function  $E(f) = \tilde{\pi}_c(f) \prod_{i=1}^n p_{\sigma^2}(f(t_i), y_i) p(t_i)$
- $f^* = SA(E, f^{(0)}, T)$

**return**  $\{(t_j^*, x_j^*)\}_{j=1}^m$ .

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# Theoretical analysis

## The topologies

The metric topology is defined by

$$d_q(f_1, f_2) = \left( \int_{t \in [0,1]} \text{dist}_M(f_1(t), f_2(t))^q p(t) dt \right)^{1/q}$$

The weak topology

$$N_\epsilon(f_0) =$$

$$\left\{ f \in \mathcal{P} : \left| \int_{[0,1] \times M} p_g p_f dt d\mu(x) - \int_{[0,1] \times M} p_g p_{f_0} dt d\mu(x) \right| \leq \epsilon, \forall g \in \mathcal{P} \right\}$$

Here  $p_f = p_{\sigma^2}(f(t), x)p(t)$ .

### Theorem

*If  $M$  is a compact Riemannian manifold and if the true underlying function  $f_0 \in \mathcal{P}$  of the regression model is Lipschitz continuous, then the posterior distribution  $\Pi(\cdot | \{(t_i, x_i)\}_{i=1}^n)$  is weakly consistent. In other words, for any  $\epsilon > 0$ ,*

$$\Pi(N_\epsilon(f_0) | \{(t_i, x_i)\}_{i=1}^n) \longrightarrow 1$$

*almost surely w.r.t. the true probability measure  $P_0^n$  as  $n \rightarrow \infty$ .*

# Theoretical analysis

## Contraction rate of DBM

### Theorem

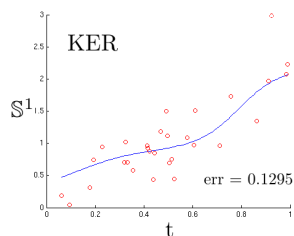
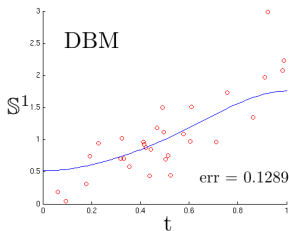
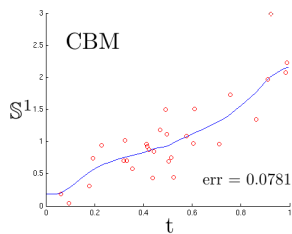
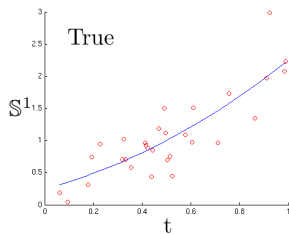
Assume  $p(t)$  is strictly positive on  $[0, 1]$ , the true function  $f_0 \in \mathcal{P}$  is Lipschitz. Assume an arbitrarily fixed  $0 < \epsilon < 1/4$  and for  $n \in \mathbb{N}$ , let  $b_n = n^{-1/2+2\epsilon}$  be the sidelength of the set  $PGF(b_n)$  and let  $\{\Pi_n\}_{n \in \mathbb{N}}$  denote the sequence of discretized BM priors on  $PGF(b_n)$ . Then there exists an absolute constant  $A_0$  and a fixed constant  $C_0$  depending only on the positive minimum value of  $p(t)$  on  $[0, 1]$ , the volume of  $M$  and the Riemannian metric of  $M$  such that  $\Pi_n(\cdot | \{(t_i, x_i)\}_{i=1}^n)$  contracts to  $f_0$  according to the rate  $\epsilon_n = \sqrt{b_n/C_0} = O(n^{-1/4+\epsilon})$ . More precisely,

$$\Pi_n(f : d_q(f, f_0) \geq A_0 \epsilon_n | \{(t_i, x_i)\}_{i=1}^n) \rightarrow 0$$

in  $P_0^n$ -probability as  $n \rightarrow \infty$ .

# Experiment

## Comparison with kernel regression



# Experiment

The hyperparameter  $c$

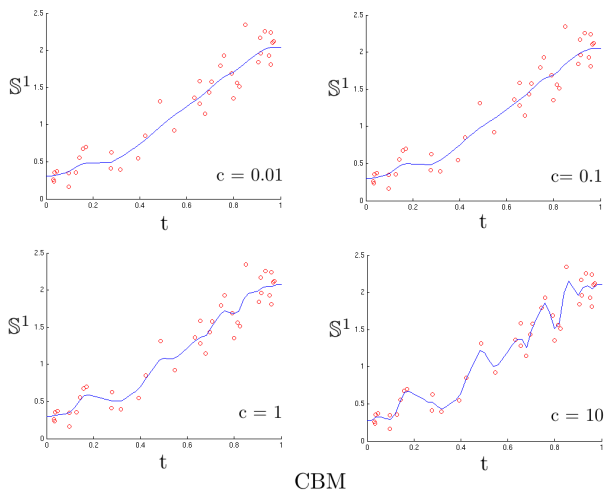


Figure : Continuous BM.



# Experiment

## The hyperparameter $c$

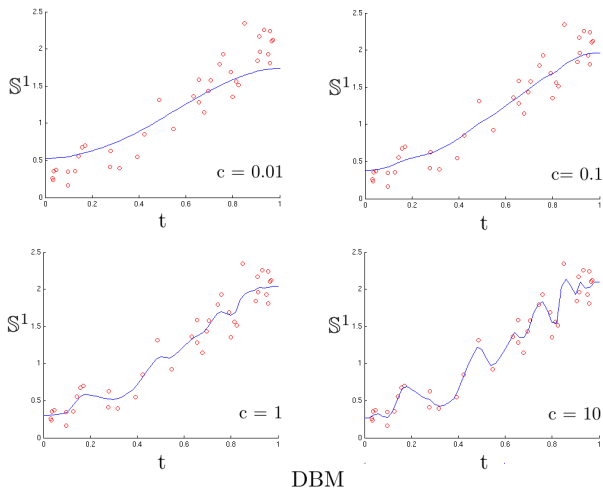


Figure : Discretized BM.

# Experiment

The sidelength parameter  $h$

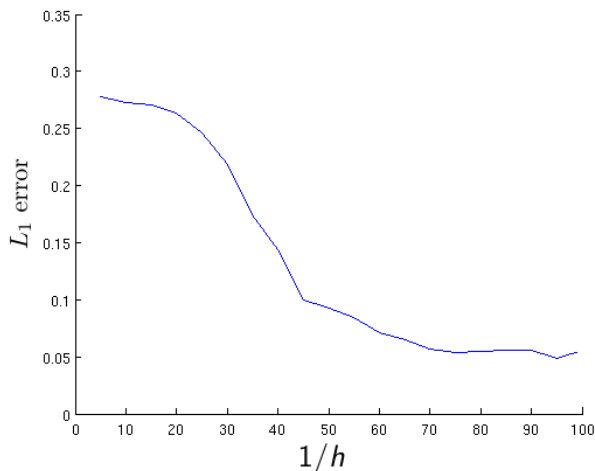


Figure :  $L_1$  error of DBM for different sidelength  $h$ .

Thank You !